Monte Carlo simulations and error analysis

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Outline of the lecture

- 1. Monte Carlo integration
- 2. Generating random numbers
- 3. The Metropolis algorithm
- 4. Monte Carlo error analysis
- 5. Cluster updates and Wang-Landau sampling
- 6. The negative sign problem in quantum Monte Carlo

1. Monte Carlo Integration

Integrating a function

• Convert the integral to a discrete sum

$$\int_{a}^{b} f(x)dx = \frac{b-a}{N} \sum_{i=1}^{N} f\left(a+i\frac{b-a}{N}\right) + O(1/N)$$

- Higher order integrators:
 - Trapezoidal rule:

$$\int_{a}^{b} f(x)dx = \frac{b-a}{N} \left(\frac{1}{2} f(a) + \sum_{i=1}^{N-1} f\left(a+i\frac{b-a}{N}\right) + \frac{1}{2} f(b) \right) + O(1/N^2)$$

• Simpson rule:

$$\int_{a}^{b} f(x)dx = \frac{b-a}{3N} \left(f(a) + \sum_{i=1}^{N-1} (3-(-1)^{i})f\left(a+i\frac{b-a}{N}\right) + f(b) \right) + O(1/N^{4})$$



High dimensional integrals

- Simpson rule with M points per dimension
 - one dimension the error is $O(M^{-4})$
 - *d* dimensions we need $N = M^d$ points the error is order $O(M^{-4}) = O(N^{-4/d})$
- An order n. scheme in I dimension is order n/d d in d dimensions!
- In a statistical mechanics model with N particles we have 6N-dimensional integrals (3N positions and 3N momenta).
- Integration becomes extremely inefficient!

Ulam: the Monte Carlo Method

- What is the probability to win in Solitaire?
 - Ulam's answer: play it 100 times, count the number of wins and you have a pretty good estimate





Throwing stones into a pond

- How can we calculate π by throwing stones?
- Take a square surrounding the area we want to measure:



• Choose *M* pairs of random numbers (*x*, *y*) and count how many points (*x*, *y*) lie in the interesting area

Monte Carlo integration

- Consider an integral $\langle f \rangle = \int_{\Omega} f(\vec{x}) d\vec{x} / \int_{\Omega} d\vec{x}$
- Instead of evaluating it at equally spaced points evaluate it at M points x_i chosen randomly in Ω:

$$\langle f \rangle \approx \frac{1}{M} \sum_{i=1}^{M} f(\vec{x}_i)$$

• The error is statistical:

$$\Delta = \sqrt{\frac{\operatorname{Var} f}{M}} \propto M^{-1/2}$$
$$\operatorname{Var} f = \langle f^2 \rangle - \langle f \rangle^2$$

• In *d*>8 dimensions Monte Carlo is better than Simpson!



- In many cases a function is large only in a tiny region
- Lots of time wasted in regions where the function is small
- The sampling error is large since the variance is large



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• Choose points not uniformly but with probability p(x):

$$\langle f \rangle = \left\langle \frac{f}{p} \right\rangle_p \coloneqq \int_{\Omega} \frac{f(\vec{x})}{p(\vec{x})} p(\vec{x}) d\vec{x} / \int_{\Omega} d\vec{x}$$

- The error is now determined by Var *f*/*p*
- Find *p* similar to *f* and such that *p*-distributed random numbers are easily available

2. Generating Random Numbers

Random numbers



http://www.idquantique.com/



Random numbers

- Real random numbers are hard to obtain
 - classical chaos (atmospheric noise)
 - quantum mechanics



http://www.idquantique.com/



Random numbers

- Real random numbers are hard to obtain
 - classical chaos (atmospheric noise)
 - quantum mechanics
- Commercial products: quantum random number generators
 - based on photons and semi-transparent mirror
 - 4 Mbit/s from a USB device, too slow for most MC simulations



http://www.idquantique.com/



• Are generated by an algorithm

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- Not random at all, but completely deterministic
- Look nearly random however when algorithm is not known and may be good enough for our purposes
- Never trust pseudo random numbers however!

Linear congruential generators

- are of the simple form $x_{n+1} = f(x_n)$
- A good choice is the GGL generator

$$x_{n+1} = (ax_n + c) \operatorname{mod} m$$

with
$$a = 16807$$
, $c = 0$, $m_{2} = 2^{3I-1}$

- quality depends sensitively on *a,c,m*.
- Periodicity is a problem with such 32-bit generators
 - The sequence repeats identically after 2³¹⁻¹ iterations
 - With 500 million numbers per second that is just 4 seconds!
 - Should not be used anymore!

Lagged Fibonacci generators

$$x_n = x_{n-p} \otimes x_{n-q} \mod m$$

- Good choices are
 - (607,273,+)
 - (2281,1252,+)
 - (9689,5502,+)
 - (44497,23463,+)
- Seed blocks usually generated by linear congruential
- Has very long periods since large block of seeds
- A very fast generator: vectorizes and pipelines very well

More advanced generators

- As well-established generators fail new tests, better and better generators get developed
 - Mersenne twister (Matsumoto & Nishimura, 1997)
 - Well generator (Panneton and L'Ecuyer, 2004)
- Based on lagged Fibonacci generators, improved with random bit shuffles
- Deep number theory enters the design of these generators



Pierre L'Ecuyer (Univ. de Montréal)

• No!

- No!
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 - Maybe?

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- Are they random enough?
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- Statistical tests for distribution and correlations
- Are these tests enough?
 - No! Your calculation could depend in a subtle way on hidden correlations!
- What is the ultimate test?
 - Run your simulation with various random number generators and compare the results

Marsaglia's diehard tests

- **Birthday spacings:** Choose random points on a large interval. The spacings between the points should be asymptotically Poisson distributed. The name is based on the birthday paradox.
- **Overlapping permutations:** Analyze sequences of five consecutive random numbers. The 120 possible orderings should occur with statistically equal probability.
- **Ranks of matrices:** Select some number of bits from some number of random numbers to form a matrix over {0,1}, then determine the rank of the matrix. Count the ranks.
- **Monkey tests:** Treat sequences of some number of bits as "words". Count the overlapping words in a stream. The number of "words" that don't appear should follow a known distribution. The name is based on the infinite monkey theorem.
- **Count the Is:** Count the I bits in each of either successive or chosen bytes. Convert the counts to "letters", and count the occurrences of five-letter "words".
- **Parking lot test:** Randomly place unit circles in a 100 x 100 square. If the circle overlaps an existing one, try again. After 12,000 tries, the number of successfully "parked" circles should follow a certain normal distribution.

Marsaglia's diehard tests (cont.)

- **Minimum distance test:** Randomly place 8,000 points in a 10,000 x 10,000 square, then find the minimum distance between the pairs. The square of this distance should be exponentially distributed with a certain mean.
- **Random spheres test:** Randomly choose 4,000 points in a cube of edge 1,000. Center a sphere on each point, whose radius is the minimum distance to another point. The smallest sphere's volume should be exponentially distributed with a certain mean.
- **The squeeze test:** Multiply 231 by random floats on [0,1) until you reach 1. Repeat this 100,000 times. The number of floats needed to reach 1 should follow a certain distribution.
- **Overlapping sums test:** Generate a long sequence of random floats on [0,1). Add sequences of 100 consecutive floats. The sums should be normally distributed with characteristic mean and sigma.
- **Runs test:** Generate a long sequence of random floats on [0,1). Count ascending and descending runs. The counts should follow a certain distribution.
- **The craps test:** Play 200,000 games of craps, counting the wins and the number of throws per game. Each count should follow a certain distribution.

Non-uniform random numbers

- we found ways to generate pseudo random numbers u in the interval [0,1]
- How do we get other uniform distributions?
 - uniform x in [a,b[: x = a+(b-a) u]
- Other distributions:
 - Inversion of integrated distribution
 - Rejection method

Non-uniform distributions

- How can we get a random number x distributed with f(x) in the interval [a,b[from a uniform random number u?
- Look at probabilities:

$$P[x < y] = \int_{a}^{b} f(t) dt =: F(y) = P[u < F(y)]$$
$$\Rightarrow x = F^{-1}(u)$$

- This method is feasible if the integral can be inverted easily
 - exponential distribution $f(x)=\lambda \exp(-\lambda x)$
 - can be obtained from uniform by $x=-1/\lambda \ln(1-u)$

Normally distributed numbers

• The normal distribution 1

$$f(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2)$$

- cannot easily be integrated in one dimension but can be easily integrated in 2 dimensions!
- We can obtain two normally distributed numbers from two uniform ones (Box-Muller method)

$$n_1 = \sqrt{-2\ln(1 - u_1)}\sin u_2$$
$$n_2 = \sqrt{-2\ln(1 - u_1)}\cos u_2$$



- Look for a simple distribution *b* that bounds $f: f(x) < \lambda h(x)$
 - Choose an *b*-distributed number *x*
 - Choose a uniform random number number $0 \le u < 1$
 - Accept x if $u < f(x)/\lambda b(x)$, otherwise reject x and get a new pair (x,u)
- Needs a good guess *h* to be efficient, numerical inversion of integral might be faster if no suitable *h* can be found



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3. The Metropolis Algorithm

Monte Carlo for classical systems

• Evaluate phase space integral by importance sampling

$$\langle A \rangle = \frac{\int_{\Omega} A(c) p(c) dc}{\int_{\Omega} p(c) dc} \longrightarrow \langle A \rangle \approx \overline{A} = \frac{1}{M} \sum_{i=1}^{M} A_{c_i}$$

• Pick configurations with the correct Boltzmann weight

$$P[c] = \frac{p(c)}{Z} = \frac{\exp(-\beta E(c))}{Z}$$

• But how do we create configurations with that distribution? The key problem in statistical mechanics!

- Metropolis Algorithm for Monte Carlo
- Simplex Method for Linear Programming

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- Krylov Subspace Iteration Methods
- The Decompositional Approach to Matrix Computations
- The Fortran Optimizing Compiler
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The Metropolis Algorithm (1953)

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Equation of State Calculations by Fast Computing Machines

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AND

EDWARD TELLER,* Department of Physics, University of Chicago, Chicago, Illinois (Received March 6, 1953)

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II. THE GENERAL METHOD FOR AN ARBITRARY POTENTIAL BETWEEN THE PARTICLES

In order to reduce the problem to a feasible size for numerical work, we can, of course, consider only a finite number of particles. This number N may be as high as several hundred. Our system consists of a square[†] con-

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Markov chain Monte Carlo

• Instead of drawing independent samples c_i we build a Markov chain

$$c_1 \rightarrow c_2 \rightarrow \dots \rightarrow c_i \rightarrow c_{i+1} \rightarrow \dots$$

- Transition probabilities $W_{x,y}$ for transition $x \rightarrow y$ need to satisfy:
 - Normalization: $\sum W_{x,y} = 1$
 - **Ergodicity:** any configuration reachable from any other

$$\forall x, y \exists n : (W^n)_{x, y} \neq 0$$

• **Balance:** the distribution should be stationary

$$0 = \frac{d}{dt}p(x) = \sum_{y} p(y)W_{y,x} - \sum_{y} p(x)W_{x,y} \Rightarrow p(x) = \sum_{y} p(y)W_{y,x}$$

• Detailed balance is sufficient but not necessary for balance

$$\frac{W_{x,y}}{W_{y,x}} = \frac{p(y)}{p(x)}$$

The Metropolis algorithm

- Teller's proposal was to use rejection sampling:
 - Propose a change with an a-priori proposal rate $A_{x,y}$
 - Accept the proposal with a probability $P_{x,y}$
 - The total transition rate is $W_{x,y} = A_{x,y} P_{x,y}$
- The choice $P_{x,y} = \min\left[1, \frac{A_{y,x}p(y)}{A_{x,y}p(x)}\right]$

satisfies detailed balance and was first proposed by Metropolis *et al*

Metropolis algorithm for the Ising model $\downarrow \uparrow \uparrow \uparrow \downarrow \downarrow$ $\uparrow \uparrow \uparrow \downarrow \downarrow \downarrow$ $\downarrow \uparrow \uparrow \uparrow \uparrow \downarrow \downarrow$ $\downarrow \uparrow \downarrow \uparrow \downarrow \downarrow$

- 1. Pick a random spin and propose to flip it
- 2. Accept the flip with probability $P = \min \left[1, e^{-(E_{new} E_{old})/T} \right]$
- 3. Perform a measurement independent of whether the proposed flip was accepted or rejected!

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Equilibration

- Starting from a random initial configuration it takes a while to reach the equilibrium distribution
- The desired equilibrium distribution is a left eigenvector with eigenvalue 1 (this is just the balance condition)

$$p(x) = \sum_{y} p(y) W_{y,x}$$

• Convergence is controlled by the second largest eigenvalue

$$p(x,t) = p(x) + O(\exp(-\lambda_2 t))$$

• We need to run the simulation for a while to equilibrate and only then start measuring

4. Monte Carlo Error Analysis

Monte Carlo error analysis

• The simple formula $\Delta A = \sqrt{\frac{\operatorname{Var} A}{M}}$

is valid only for independent samples

• The Metropolis algorithm gives us correlated samples! The number of independent samples is reduced

$$\Delta A = \sqrt{\frac{\operatorname{Var} A}{M} \left(1 + 2\tau_A\right)}$$

• The autocorrelation time is defined by

$$\tau_{A} = \frac{\sum_{t=1}^{\infty} \left(\left\langle A_{i+t} A_{i} \right\rangle - \left\langle A \right\rangle^{2} \right)}{\operatorname{Var} A}$$

Binning analysis

• Take averages of consecutive measurements: averages become less correlated and naive error estimates converge to real error



0

0

2

binning level l

8

10

a smart implementation needs only O(log(N)) memory for N measurements

Seeing convergence in ALPS

- Look at the ALPS output in the first hands-on session
- 48 x 48 Ising model at the critical point
 - local updates:

Name	Count	Mean	Error	Tau	Method
Susceptibility	52529	401.08	11.3 not converged	99.1	binning

• cluster updates:

Name	Count	Mean	Error	Tau	Method
Susceptibility	113433	421.642	1.57	0.821	binning

Correlated quantities

• How do we calculate the errors of functions of correlated measurements?

• specific heat
$$c_V = \frac{\langle E^2 \rangle - \langle E \rangle^2}{T^2}$$

• Binder cumulant ratio $U = \frac{\langle m^4 \rangle}{\langle m^2 \rangle^2}$

- The naïve way of assuming uncorrelated errors is wrong!
- It is not even enough to calculate all crosscorrelations due to nonlinearities except if the errors are tiny!









$$\left\langle U \right\rangle \approx \overline{U} = \frac{1}{M} \sum_{i=1}^{M} U_{i}$$

$$\Delta U \approx \sqrt{\frac{1}{M(M-1)} \sum_{i=1}^{M} \left(U_i - \overline{U} \right)^2}$$

Simplest idea: split the time series and evaluate for each segment



Problem: can be unstable and noisy for nonlinear functions such as X/Y

Jackknife-analysis

Jackknife-analysis

$$U_0 = \frac{1}{M} \sum_{i=1}^M f(X_i, Y_i) \frac{f(X_1, Y_1) - f(X_2, Y_2) - f(X_3, Y_3)}{f(X_1, Y_1) - f(X_2, Y_2) - f(X_3, Y_3) - \dots - f(X_M, Y_M)}$$

Jackknife-analysis



Jackknife-analysis



Jackknife-analysis



$$\langle U \rangle \approx U_0 - (M-1)(\overline{U} - U_0)$$
 $\overline{U} = \frac{1}{M} \sum_{i=1}^M U_i$

$$\Delta U \approx \sqrt{\frac{M-1}{M} \sum_{i=1}^{M} \left(U_i - \overline{U} \right)^2}$$

ALPS.Alea library

- The ALPS class library implements reliable error analysis
 - Adding a measurement:

```
alps::RealObservable mag;
...
mag << new_value;</pre>
```

• Evaluating measurements

```
std::cout << mag.mean() << " +/- " << mag.error();
std::cout "Autocorrelation time: " << mag.tau();</pre>
```

- Correlated quantities?
 - Such as in Binder cumulant ratios

```
\left< m^4 \right> / \left< m^2 \right>^2
```

• ALPS library uses jackknife analysis to get correct errors

alps::RealObsEvaluator binder = mag4/(mag2*mag2);
std::cout << binder.mean() << " +/- " << binder.error();</pre>

 Critical slowing down, cluster updates and Wang-Landau sampling

Autocorrelation effects

• The Metropolis algorithm creates a Markov chain

$$c_1 \rightarrow c_2 \rightarrow \dots \rightarrow c_i \rightarrow c_{i+1} \rightarrow \dots$$

• successive configurations are correlated, leading to an increased statistical error

$$\Delta A = \sqrt{\left\langle \left(\overline{A} - \left\langle A \right\rangle\right)^2 \right\rangle} = \sqrt{\frac{\operatorname{Var} A}{M} (1 + 2\tau_A)}$$

- Critical slowing down at second order phase transition $\tau \propto L^2$
- Exponential tunneling problem. at first order phase transition $\tau \propto \exp(L^{d-1})$

From local to cluster updates

- Energy of configurations in Ising model
 - -J if parallel: $\uparrow \uparrow \downarrow \downarrow$
 - + *J* if anti-parallel: $\downarrow \uparrow \downarrow \uparrow$
- Probability for flip
 - Anti-parallel: flipping lowers energy, always accepted $(\downarrow) \uparrow \qquad \uparrow \qquad \Delta E = -2J \Rightarrow P = \min(1, e^{-2\Delta E/T}) = 1$
 - Parallel:

 $(\uparrow) \uparrow \longrightarrow \downarrow \uparrow \qquad \Delta E = +2J \Rightarrow P = \min(1, e^{-2\Delta E/T}) = \exp(-2\beta J)$

no change with probability $1 - \exp(-2\beta J)$!!!

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Alternative: flip both!

$$\begin{array}{c} \uparrow \uparrow \\ \uparrow \\ \uparrow \\ \uparrow \\ \uparrow \\ \uparrow \\ \downarrow \\ \end{array} \begin{array}{c} P = \exp(-2J/T) \\ P = 1 - \exp(-2J/T) \\ P = 1 - \exp(-2J/T) \\ \end{array}$$
- No critical slowing down (Swendsen and Wang, 1987) !!!
- Ask for each spin: "do we want to flip it against its neighbor?"
 - antiparallel: yes
 - parallel: costs energy
 - Accept with

$$P = \exp(-2\beta J)$$

- $P = 1 \exp(-2\beta J)$
- Otherwise: also flip neighbor!
- Repeat for all flipped spins => cluster updates

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$$\downarrow \uparrow ? \uparrow ? \uparrow \downarrow \downarrow \downarrow$$

Shall we flip neighbor?
$$\uparrow ? \uparrow \downarrow \uparrow \downarrow \downarrow \downarrow \downarrow$$

$$\downarrow \uparrow \downarrow \downarrow \uparrow \downarrow \uparrow \uparrow \uparrow \uparrow \uparrow$$

$$\uparrow \uparrow \downarrow \downarrow \uparrow \uparrow \uparrow \uparrow \uparrow$$

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Done building cluster Flip all spins in cluster

The loop algorithm (Evertz et al, 1993)

- Generalization of cluster updates to quantum systems
- Loop-like clusters update world lines of spins



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First order phase transitions

- Tunneling problem at a first order phase transition is solved by *changing the ensemble*. to create a flat energy landscape
 - Multicanonical sampling (Berg and Neuhaus, Phys. Rev. Lett. 1992)
 - Wang-Landau sampling (Wang and Landau, Phys. Rev. Lett. 2001)
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"flat-histogram" weight



$$n_w(E) = \frac{1/g(E)}{2} \cdot g(E)$$

"flat-histogram" weight

How do we obtain the weights?



$$n_w(E) = \frac{1/g(E)}{f} \cdot g(E)$$

"flat-histogram" weight

How do we obtain the weights?



Flat-histogram MC algorithms

- Multicanonical recursions
 B. A. Berg and T. Neuhaus (1992)
- ► Wang-Landau algorithm F. Wang and D.P. Landau (2001)
- Quantum version M. Troyer, S. Wessel and F. Alet (2003)

The Wang-Landau algorithm

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• Start with "any" ensemble

$$w(E) = \frac{1}{\tilde{g}(E)} \qquad \tilde{g}(E) = 1$$

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estimated density of $\tilde{g}(E) = 1$ states

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 $\tilde{g}(E) = 1$ states

estimated

1 .

• Simulate using Metropolis algorithm

 \mathcal{U}

$$p(E_1 \to E_2) = \min\left(1, \frac{w(E_2)}{w(E_1)}\right) = \min\left(1, \frac{\tilde{g}(E_1)}{\tilde{g}(E_2)}\right)$$

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estimated

• Reduce modification factor f when histogram is flat.

Wang-Landau in action Movie by Emanuel Gull (2004)

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6. The negative sign problem in quantum Monte Carlo

• Not as easy as classical Monte Carlo

$$Z = \sum_{c} e^{-E_{c}/k_{B}T}$$

- Calculating the eigenvalues E_c is equivalent to solving the problem
- Need to find a mapping of the quantum partition function to a classical problem

$$Z = \operatorname{Tr} e^{-\beta H} \equiv \sum_{c} p_{c}$$

• "Negative sign" problem if some $p_c < 0$

- Feynman (1953) lays foundation for quantum Monte Carlo
- Map quantum system to classical world lines

The Physical Review

A journal of experimental and theoretical physics established by E. L. Nichols in 1893

Second Series, Vol. 91, No. 6

SEPTEMBER 15, 1953

Atomic Theory of the λ Transition in Helium

R. P. FEYNMAN California Institute of Technology, Pasadena, California (Received May 15, 1953)

- Feynman (1953) lays foundation for quantum Monte Carlo
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Use Metropolis algorithm to update world lines
The negative sign problem

• In mapping of quantum to classical system

$$Z = \mathrm{Tr}e^{-\beta H} = \sum_{i} p_{i}$$

- there is a "sign problem" if some of the $p_i < 0$
 - Appears e.g. in simulation of electrons when two electrons exchange places (Pauli principle)



The negative sign problem

• Sample with respect to absolute values of the weights

$$\langle A \rangle = \sum_{i} A_{i} p_{i} / \sum_{i} p_{i} = \frac{\sum_{i} A_{i} \operatorname{sgn} p_{i} |p_{i}| / \sum_{i} |p_{i}|}{\sum_{i} \operatorname{sgn} p_{i} |p_{i}| / \sum_{i} |p_{i}|} = \frac{\langle A \cdot \operatorname{sign} \rangle_{|p|}}{\langle \operatorname{sign} \rangle_{|p|}}$$

• Exponentially growing cancellation in the sign

$$\langle sign \rangle = \frac{\sum_i p_i}{\sum_i |p_i|} = Z/Z_{|p|} = e^{-\beta V(f - f_{|p|})}$$

• Exponential growth of errors

$$\frac{\Delta sign}{\langle sign \rangle} = \frac{\sqrt{\langle sign^2 \rangle - \langle sign \rangle^2}}{\sqrt{M} \langle sign \rangle} \approx \frac{e^{\beta V (f - f_{|p|})}}{\sqrt{M}}$$

• NP-hard problem (no general solution) [Troyer and Wiese, PRL 2005]

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- We simulate a ferromagnet and expect to learn something about a spin glass?
 - This is the idea behind the proof of NP-hardness

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 - Bosonic atoms in optical lattices
 - Helium-4 supersolids
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- 4. Use approximate methods
 - Dynamical mean field theory (DMFT)

7. Diverging Length Scales and Finite Size Scaling

Divergence of the correlation length $\boldsymbol{\xi}$

• Typical length scale ξ divegres at phase transition at Tc



• To avoid system size effects we need to have $L >> \xi \rightarrow \infty$

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Renormalization group and scaling

- As the length scale ξ diverges, "microscopic details" can be ignored
 - Physics happens at "large" length scale ξ
 - Microscopic length scale a of lattice can be ignored
 - All models with same symmetry converge to the same fixed point

- Fixed point is scale free
 - The only length scale ξ diverges
 - Self-similarity and fractal behavior
 - Power laws are scale free functions

 $m \propto (T_c - T)^{\beta}$

 $\xi \propto |T - T_c|^{\nu}$

"Finite-size scaling"

• Infinite system

$$M \propto \left(T_c - T\right)^{\beta} \qquad \qquad \xi \propto \left(T_c - T\right)^{-\nu}$$

write M in terms of length scale $\boldsymbol{\xi}$

$$\Rightarrow M(T) = M(\xi) \propto \xi^{-\beta/\nu}$$

• finite systems: L acts as cutoff to ξ

$$M(T,L) = M(\xi,L) = M(\xi/L) \propto \begin{cases} \xi^{-\beta/\nu} & L >> \xi \\ L^{-\beta/\nu} & L << \xi \end{cases}$$

• We can obtain critical exponents β , v from finite size effects



A quantum antiferromagnet

- Quantum phase transition in a 2D Heisenberg antiferromagnet
- Susceptibility

 $\chi_s \propto L^{2-\eta}$

• Structure factor of magnetization

 $S(Q) = L^2 m \propto L^{2-z-\eta}$

- Scaling fits give z and η
- Additional dynamical critical exponent z is only difference from classical FSS

 $\xi_{\tau} \propto \xi^z$

